On Public Opinion Polls and Voters' Turnout*

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Abstract:

This paper studies the effects that the revelation of information on the electorate's

preferences has on voters' turnout decisions. The experimental data show that closeness

in the division of preferences induces a significant increase in turnout. Moreover, for

closely divided electorates (and only for these electorates) the provision of information

significantly raises the participation of subjects supporting the slightly larger team rela-

tive to the smaller team. We show that the heterogeneous effect of information on the

participation of subjects in different teams is driven by the subjects' (incorrect) beliefs

of casting a pivotal vote. Simply put, subjects overestimate the probability of casting a

pivotal vote when they belong to the team with a slight majority, and choose the strat-

egy that maximizes their utility based on their inflated probability assessment. Empirical

evidence on gubernatorial elections in the U.S. between 1990 and 2005 is consistent with

our main experimental result. Namely, we observe that the difference in the actual vote

tally between the party leading according to the polls and the other party is larger than

the one predicted by the polls only in closely divided electorates. We provide a behavioral

model that explains the main findings of our experimental and empirical analyses.

KEYWORDS: Voter Turnout, Public Opinion Polls, Experimental Economics.

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1. Introduction

In large electorates the probability of casting a pivotal vote is close to zero regardless of the actual distribution of preferences. A poll pointing to an evenly split electorate, however, may affect the voters' beliefs on the probability of casting a pivotal vote and, therefore, the voters' turnout decisions. Indeed, a lively debate is being carried in several countries on whether or not polls affect electoral results. A fundamental difficulty when trying to empirically assess the causal effect of public opinion polls on the individuals' turnout decisions is that of omitted variables. Several factors, like valence characteristics of candidates and their chosen platforms, affect not only individuals' turnout but also the public opinion polls.

This paper analyses experimentally and empirically the impact that the provision of information on the electorate's distribution of preferences has on the voters' participation decisions.² Our experiment compares the subjects' participation decisions in an election when they know the exact distribution of preferences of the electorate to their decisions when they only know their own preferences. Our objective is to uncover any behavioral effects that the provision of information may have on the voters. Additionally, we collected, through a survey administered at the beginning and at the end of our experiment, the

¹In the last presidential elections in the U.S., for example, individual voters that supported Ralph Nader and resided in states where the election was predicted to be close traded their votes with John Kerry's supporters that lived in states where the election was expected to be lopsided in favor of one candidate. People that traded votes felt that now their vote "really counted." As related in votepair.org/stories: "I live in Utah. The most republican state in the nation. I happen to be a democrat who voted for Gore. My vote did not count because of the stupid electoral college. By swapping my vote, I can finally have my vote count for a democrat."

²See Goeree and Großer (2005) and Taylor and Yildirim (2005) for recent theoretical studies of the effects of information on the electorate's behavior.

subjects' estimated probabilities of casting a pivotal vote for all the different distributions of preferences. This allow us to assess whether the subjects' behavior is a consequence of their beliefs or despite thereof.

The experimental results show that closeness in the division of preferences induces a significant increase in turnout. Perhaps more surprisingly, in closely divided electorates (and only for these electorates) the provision of information significantly raises the participation of subjects supporting the slightly larger team relative to the smaller team –we refer to this behavior as the bandwagon effect of polls. This behavior contradicts the qualitative predictions of the unique quasi-symmetric equilibrium of the theoretical model underlying the experiments. According to the equilibrium conditions the provision of information on the electorate's preferences should induce voters in the majority to participate less frequently because they free ride on the voting of other individuals supporting the same alternative. At the same time polls should stimulate the participation of voters in the minority to offset the advantage of the other alternative. These requirements of the mixed strategy Nash equilibrium seem counter-intuitive and are not supported by the experimental data.³

To uncover the root causes behind the bandwagon effect of polls we incorporate into the analysis the subjects' responses to the surveys. This analysis shows that the heterogeneous effect of information on the participation of subjects in different teams is driven by the subjects' (incorrect) beliefs of casting a pivotal vote. Simply put, subjects overestimate the

³A similar behavioral departure from mixed strategies Nash equilibrium was documented by Rapaport et al. (2002) in an experimental study of market entry with asymmetric players.

probability of casting a pivotal vote when they belong to the team with a slight majority, and choose the strategy that maximizes their utility based on their inflated probability assessment. This conjecture was first formalized by Riker and Ordeshook (1967). To the best of our knowledge, the current paper is the first attempt to formally test this hypothesis.

The observed bandwagon effect of polls is consistent with previous experimental studies. While studying the incidence of reform in the presence of individual-specific uncertainty, Cason and Miu (2005) find that the participation rates of the majority are higher than the participation rates of the minority. In an independent study, Großer et al. (2005) examine the welfare implications of endogenous voter participation using a different experimental design that includes floating voters. They also find that the majority participates more than the minority but this difference is not statistically significant. Their experimental design allows them to test this hypothesis only using the electorate as the unit of observation. Our experiment, like Cason and Miu's (2005), is especially designed to use the subjects as our unit of observation, granting us the possibility to differentiate between distributions with enough observations for each one to be able to perform statistically meaningful tests.

Finally, Levine and Palfrey (2005) experimentally test the predictions of Palfrey and Rosenthal (1985) whereby participation costs are heterogeneous and privately known. They find that subjects in the small team vote with higher frequency than subjects in the large team. Unlike our experimental design, theirs doesn't directly test for the effects of the provision of information. Perhaps more importantly, their study [as well as that of

Großer et al. (2005)] reveals the exact vote tally at the end of each round thus allowing the subjects to gain experience and learn over rounds. This was done in order to check whether the subjects' behavior converges with experience to the one predicted by the pure strategy equilibrium of the game they analyzed.

This paper's objective is not to examine the predictions of a particular model, but rather to reveal the individuals' behavioral reactions to the publication of public opinion polls. Therefore, we use a random and anonymous reassignment procedure specifically to reduce repeated game incentives and minimize the effects of learning. We believe this is the right experimental design given the objectives of our study.

We test the external validity of our main experimental result using a newly culled data set on gubernatorial elections in the U.S. since 1990. For these purposes we use as our proxy to closeness in the distribution of preferences the results of pre-election polls published by newspapers within one week before the elections. The observed evidence is consistent with the experimental results. Namely, in elections where the polls pointed to a narrow margin between the two parties the difference in the elections' vote share of the two parties is greater than the difference predicted by the polls. The effect above is not present in electoral contests that were expected to be lopsided according to the polls' predictions.

The individuals' behavioral pattern has significant implications in two different contexts. The immediate one is in the context of voting, where our results point to an interesting behavioral phenomenon that has been overlooked by the related literature. This behavior has important implications on the widespread policy debate on the desirability

of publishing polls close to an election date. On the one hand, supporters of the ban claim that the observed inclination of people to vote for candidates leading in the surveys may lead to the manipulation of polls before elections by parties with vested interests. On the other hand, opponents to the measure claim that a ban on polls before elections suppresses the freedom of expression.⁴ Experimentally, we show that the bandwagon effect is a direct consequence of higher voter participation and not necessarily of voters changing their preferences. This suggests that a policy geared to increase voters' participation can substantially offset the effects of polls.

The observed phenomenon has also broader implications regarding the empirical relevance of mixed strategy equilibria in more general setups. When a player can choose between two alternatives she may use a mixed strategy only when she is indifferent between the two. In an asymmetric environment as the one proposed here, players in the small team would be indifferent between voting and abstaining only if players in the large team vote with a relatively lower frequency than players of the small team. Moreover, each player should have the correct beliefs regarding the mixed strategies used by the rest of the players. The equilibrium strategies (and beliefs) are not necessarily intuitive, especially in setups with only slight differences between the players. Thus, for these conditions to hold behaviorally probably requires that the game is played with a considerable amount of repetition to facilitate experience and learning.

We propose an alternative theoretical explanation that relaxes Nash equilibrium but is

⁴According to the Foundation for Information/ESOMAR a ban on the publication of opinion polls exists in 30 out of 66 countries surveyed in their study published in 2003. Nowadays, a lively debate is being conducted in several countries, like Canada, France, Ireland, The Philippines and Russia.

consistent with the voters' beliefs and behavior observed in the laboratory. In particular, we show that if individuals believe that in a close election the probability of voting is sufficiently high and similar for every voter regardless of team sizes, then optimal behavior with respect to these beliefs gives rise to voting patterns consistent with the ones observed in the current study. While these beliefs cannot be part of equilibrium with groups of unequal sizes, they are consistent with the documented departures from quasi-symmetric equilibrium strategies in other contexts as well.

The paper is organized as follows. The next section presents the theoretical framework underlying our experiment. A detailed description of our experimental design appears in Section 3. Section 4 shows the main experimental results of the paper. Section 5 test the external validity of our experimental results. We present an alternative theoretical explanation for the subjects behavior in Section 6. The last section of the paper concludes. The proof of our theoretical result appears in the appendix.

2. Theoretical Framework

The theoretical framework we consider is based on the seminal contribution of Palfrey and Rosenthal (1983). There are n risk neutral individuals ($n \geq 3$). Individuals have to decide between two alternatives $\{A, B\}$. The alternative is chosen via simple plurality rule; that is, the alternative with the greater number of votes is chosen. In the event of a tie each alternative is selected with equal probability. This is a collective choice problem: the chosen alternative applies to all the individuals.

Each individual has preferences over the two alternatives. Let V denote the utility difference to an individual between the event that her favored alternative is elected and the event that the other alternative wins the election. Each individual has to decide whether to vote or abstain.⁵ Let us denote by s_i the strategy of individual i (let $s_i = 1$ when individual i votes and $s_i = 0$ otherwise). All the individuals make their strategy choices simultaneously. There is a positive cost C > 0 associated with the act of voting. V and C are common knowledge and identical to all the individuals. We assume that V > 2C.

In this setup, a rational individual votes if and only if

$$V \times P(1, s_{j \neq i}) - C \ge V \times P(0, s_{j \neq i}),$$

where P denotes the probability that individual i's preferred alternative is chosen and $s_{j\neq i}$ is a profile that describes the strategy of all the individuals excluding individual i.

Clearly, a rational individual participates in the election only if, given the other individuals' strategies, her participation affects the probability that her preferred alternative is chosen. In other words, an individual may turn out to vote only when she is pivotal.

We analyze the game above under two different frameworks regarding the individuals' information about the distribution of preferences. The first scenario focuses on a symmetric private value model of voting. Accordingly, each voter knows the alternative that she

⁵In the present framework voting against one's preferred alternative is strictly dominated by not voting. Therefore, we rule out this possibility and, whenever we say that an individual votes, we imply that she is voting in support of her preferred alternative.

favors and that the probability that any other individual prefers any given alternative is the same for both alternatives. The individuals' probability distributions are stochastically independent.

We focus on symmetric Bayesian Nash Equilibrium (BNE) as the relevant equilibrium concept for the symmetric private value model of voting. This equilibrium concept assumes that the every individual's decision to participate is independent of the alternative that she favors because of the symmetric common prior over the individuals' distribution of preferences; that is, all the individuals randomize between voting for their preferred alternative and abstaining with the same probability. (A formal definition appears in Appendix A.)

In the second scenario the number of voters favoring each alternative is commonly known. This is exactly the framework analyzed by Palfrey and Rosenthal (1983). This complete information game has multiple Nash equilibria. The solution concept that generates unique predictions for the game is that of totally quasi-symmetric mixed strategy Nash equilibrium (QSNE). According to this equilibrium concept all the individuals supporting the same alternative are using the same strategy. Moreover, this strategy involves voting with a probability strictly between zero and one. Note that, unlike the BNE, in the QSNE individuals supporting different alternatives are not necessarily mixing with the same probability. (See Appendix A for a formal definition of this equilibrium concept.)

For the purposes of our experimental study we focus on electorates of seven individuals and set V = 10 and C = 4. We choose an odd number of participants in each electorate to rule out equilibrium in pure strategies (except for the case where all the participants

share the same preferences). When the subjects know the distribution of preferences symmetric equilibria do not exist for this configuration either. In fact, with seven players and two alternatives there exists a unique totally mixed QSNE and a unique totally mixed symmetric BNE. Table 1 provides the point predictions for the unique BNE and QSNE.

[Table 1 about here]

Note that in the QSNE for every distribution of preferences individuals in the minority vote with higher probability than individuals in the majority. This result is a direct consequence of the mixed strategies equilibrium's requirement that individuals' should be indifferent between voting and abstaining. Since individuals would be willing to vote only if the probability of casting a pivotal vote is positive, they have to expect that with a high enough probability the number of votes in support for each team would be equal, or differ by only one vote. To satisfy that requirement individuals supporting the large team should vote with lower probability than individuals supporting the small team.

3. Experimental Design

The experiment was run at the RatioLab - The Center for Rationality and Interactive Decision Theory at The Hebrew University of Jerusalem. The 84 subjects in this experiment were recruited from the pool of undergraduate and graduate students from The Hebrew University and had no previous experience in experiments related to voters' participation.

In each session 21 subjects participated as voters. The experiments were conducted

via computers. Before the experiment started an experimental administrator read the instructions aloud. We also asked several hypothetical questions at the end of the instructions to check subjects' comprehension of the procedure (the instructions and the questionnaire are located in Appendix B). The experiment began after all subjects had solved all questions successfully. The experiment lasted for about ninety minutes. Each subject received 80 tokens as a participation fee and subsequent earnings according to the payoffs specified in the experiment. Average earnings were equal to 244 tokens. We converted each token to NIS 0.25 and paid the subjects in cash in private at the end of the session.⁶ Throughout the experiment we ensured anonymity and effectively isolated each subject in a cubicle to minimize any interpersonal influence that could stimulate uniformity of behavior. Communication among subjects was not allowed throughout the session.

Each experimental session entailed 20 independent rounds. In each round we randomly divided 21 subjects into three electorates of seven participants each. At the beginning of each round an equal probability rule randomly assigned each subject to one of two teams: Green or Blue. A subject earns 10 tokens if the team she prefers is selected by majority voting in an election. Voting entails a cost of 4 tokens.

The sequence of events is as follows. Subjects know that the round is divided into two stages, and that each subject will decide whether to vote or abstain in each stage. Every subject knows that her decision in one stage is independent from her decision in

⁶That is, subjects on average earned NIS 61 for roughly 90 minutes of their time. The hourly minimum wage in Israel is slightly below NIS 20. The current exchange rate is NIS 4.5 per U.S. dollar.

the other stage. In the first stage of each round each subject knows only her preferred color. She decides whether to vote or abstain. After all the participants make their decisions we proceed to the second stage of the round. In this stage subjects are told the electorate's distribution of preferences. Note that subjects don't receive any information on the subjects' participation decisions in the round's first stage. Subjects have to decide again whether or not to vote. After all the subjects choose an action, they learn the selected teams of the first and second elections, their corresponding payoffs for the round, and their cumulative payoffs – no information is provided on the number of subjects that voted for a given team. Ties are always broken by an equal probability rule. At the end of each round subjects are randomly reshuffled between electorates and each subject's preferred color is again randomly decided.

In addition to playing this game each subject completed a survey that asked her to assess the probability of casting a pivotal vote for each possible team size. Every subject completed the same survey twice – before the beginning of the first round and after finishing the last round.

4. Experimental Results

This section presents the effects of revealing information about the electorate's distribution of preferences on the subjects' turnout decisions. To clarify the exposition we divide this section into two subsections. The first subsection presents the basic results on the impact of information provision on subjects' participation decisions. The second subsection reports

the results taking into account not only subjects' actions but also their beliefs.

For all the tests reported below the unit of observation is the subject. For the nonparametric tests we consider, for each subject, the average across all the different rounds. This eliminates possible correlations across repeated observations of a given subject. Therefore, the statistics reported are averages of the subjects' averages. In the regression analysis, however, we use all the available data, adopting a random effects specification with the subject as the random factor.

4.1. The Effect of Information on Subjects' Turnout Decisions

Figure 1 depicts the average turnout rate before the provision of information and the average turnout rate after information is revealed, as a function of the different distribution of preferences. The figure also includes the equilibrium's predicted turnout rate.

[Figure 1 about here]

The figure clearly indicates that closeness in the division of preferences induces a significant increase in turnout. Whereas the average turnout rate before the provision of information is slightly below 25 percent, the average turnout rate for a distribution of teams of sizes three and four is 40 percent (the difference between the two is statistically significant with z = 3.125, p < 0.001, two-sided sign test using the normal approximation to the binomial distribution).

The provision of information for other divisions of the electorates doesn't have a signif-

icant impact on the subjects' turnout relative to their turnout rates before the provision of information (p > 0.8 when the division of teams is five versus two; p > 0.65 when the division is six versus one; and p > 0.8 when the division is seven versus zero, all according to a two-sided sign test). Moreover, the observed rates aren't substantially different from the equilibrium's prediction. The turnout rate is higher than the equilibrium's prediction for distributions of seven versus zero and five versus two, whereas turnout is lower than the equilibrium's prediction for a distribution of six versus one. For closely divided preferences, on the contrary, we observe important quantitative differences between the subjects' turnout and the predictions of the theoretical model underlying the experiment.

Although Figure 1 reveals a clear and significant effect of closeness on participation, the figure masks important and unexpected differences between teams for a given distribution of preferences. The heterogeneous effect of closeness between teams is presented in Figure 2, which decomposes turnout as a function of the size of the teams. Note that a team of size j implies that the distribution of the electorate's preferences is (j, 7 - j).

This figure shows the most startling effect that emerges from our experiment: For closely divided electorates the effect of information on voter participation is not homogenous across teams of different sizes. In particular, the provision of information significantly raises the participation of voters supporting the slightly larger team relative to the participation of voters supporting the smaller team, thus affecting the election's results.

⁷The correlation between closeness and turnout observed in the laboratory is consistent with results in the related empirical literature [see, for example, Shachar and Nalebuff (1999)]. Feddersen and Sandroni (2006) show that this correlation can be explained using a model where voters have ethical preferences. Coate and Conlin (2004) provide empirical evidence supporting the ethical voters approach.

[Figure 2 about here]

In other words, for closely divided electorates revealing information on subjects' preferences causes an important increase on the participation of all the subjects: Subjects that belong to teams of size three and four vote more often after learning the distribution of preferences. The effect, however, is stronger for subjects that belong to the slightly larger team. The turnout rate for subjects that belong to a team with four supporters is more than twenty percent higher than the turnout rate of subjects that belong to a team of three. This behavior contradicts the quantitative and qualitative predictions of the theoretical model. Accordingly, members of the minority should vote with a higher probability than the members of the majority to offset the advantage of the majority. Moreover, the provision of information should induce a decrease in the turnout rate of the majority because of free riding of its members.

We don't observe a similar effect for electorates with a more lopsided division of preferences. For electorates that aren't closely divided, revealing the distribution of preferences doesn't affect the turnout rate of subjects supporting the small team but lowers the turnout of subjects supporting the large team. For example, we see an important decrease in the participation of subjects after learning that they belong to a team of size seven. A similar phenomenon occurs for subjects that belong to a team of six subjects. Note that these subjects turn out in a frequency lower than the frequency of a subject that is the sole supporter of an alternative. An analogous situation occurs when the subjects' preferences

are divided between teams of five and two members. This behavior, which seems to be a consequence of free riding, is in accordance with the equilibrium's predictions.

The different effect of closeness on subjects conditional on the size of the team they support is evident from the estimation of the following participation equation:

$$Vote_Inf_{i,t} = 1 \{ \beta_0 + \beta_1 Vote_NoInf_{i,t} + \beta_2 Majority_{i,t} + \beta_3 round_t + \alpha_i \ge 0 \}$$
 (4.1)

where $1\{\cdot\}$ is an indicator function that takes the value 1 if the left hand side of the inequality inside the brackets is greater than or equal to zero, and 0 otherwise; $Vote_Inf_{i,t}$ reflects subject i's participation decision in the second stage of round t after the provision of information on the distribution of preferences. The covariates account for the subject's decision in the first stage of round t before the provision of information ($Vote_NoInf_{i,t}$), and whether or not the subject belongs to the large team in an electoral contest ($Majority_{i,t}$). We also include in the analysis a time trend ($round_t$) to capture the fact that subjects may systematically change their strategy as a consequence of learning from round to round; and a subject specific constant effect (α_i) that captures random disturbances (constant through time) that characterize subject i.

We estimate equation (4.1) separately for each different distribution of subjects' preferences using a random effects probit estimation.⁸ Table 2 presents the estimated coefficients.

⁸A similar estimation strategy was used in an experimental context by Frechette et al. (2005). That study's main focus is the analysis of the impact of open versus closed amendment rules in models of bargaining.

[Table 2 about here]

The table quantifies the most striking of our results: When the electorate is closely divided, the provision of information on subjects' preferences significantly raises the participation of subjects in the majority relative to the minority. The team size effect when the electorate is divided into teams of three versus four subjects is positive, large in value relative to the other coefficients and statistically significant. It increases the probability of voting by slightly over 10 percent for the average subject. We also observe a significant negative effect of rounds, pointing to a learning process that induces subjects to reduce their participation in elections.⁹ Interestingly, subjects' participation decision in the first stage of each round doesn't explain their actions after the provision of information.

The subjects' behavior is qualitatively different when the difference in the number of supporters for each team is relatively large. When the difference in the number of supporters is of three or five subjects the provision of information doesn't affect the participation of subjects in the majority any differently than it affects the participation of subjects in the minority. For these groups compositions, moreover, the number of rounds elapsed doesn't affect participation.

Contrary to closely divided groups, when the difference in the number of supporters for each team is relatively large the best predictor for subjects' participation decisions

⁹We test the same model including the interaction between majority and number of rounds as an additional covariate. The coefficient for this covariate is not significant; thus, the significant differences between majority and minority don't disappear over rounds.

after the provision of information is the subjects' actions in the first stage of each round. That is, there are subjects that reveal a preference for participation in the first stage, and therefore these subjects are the ones turning out to vote in the second stage when the electorate's preferences are not closely divided. This seems to be particularly the case in very lopsided contests (6 versus 1) where the coefficient of the first stage decision is not only highly statistically significant, but also large in value relative to the other coefficients. The probability of voting in the second stage is 30 percent higher for the average subject that belongs to a team of size one and voted in the first stage relative to the average subject that didn't vote in the first stage. The marginal effect of voting in the first stage on the probability of voting in the second stage is 24 percent for subjects in teams of size 6.

The fact that the provision of information significantly raises the participation of subjects in the majority relative to the minority in closely divided electorates not only contradicts the intuitions behind the mixed strategies Nash equilibrium concept but also those of the alternative quantal response equilibrium concept. Interestingly, however, the quantal response equilibrium predicts that the observed participation rate should be higher (lower) than that predicted by the Nash equilibrium when the participation rate predicted by the Nash equilibrium is below (above) 0.5 (Goeree and Holt, 2005). This prediction is borne out by the data both before and after the provision of information. Moreover, this prediction finds additional support in the next section where we analyze the subjects' beliefs.

4.2. The Effect of Subjects' Beliefs on Their Turnout Decisions

This subsection incorporates into the analysis the surveys' answers to better account for the subjects' strategies. As already pointed out, these surveys, conducted at the beginning and at the end of the experiment, asked every subject to quantify the probability of casting a pivotal vote for every possible distribution of preferences. Theoretically, the equilibrium probability of casting a pivotal vote depends only on 2C/V —the voting cost divided by half the benefits of a victory of the subject's preferred alternative. Given that we hold both constant, the equilibrium's predicted probability of casting a pivotal vote is constant as well regardless of the distribution of preferences. (In our application with a benefit of 10 tokens and a cost of 4 tokens this probability is equal to 0.8).

Figure 3 depicts the average subjects' beliefs of casting a pivotal vote as a function of the size of the team. The figure includes the results of the survey taken at the beginning (labeled survey 1 in the figure) and at the end of the experiment (survey 2). This figure also includes the frequencies of elections in which at least one subject was pivotal based on the other subjects actual behavior.

[Figure 3 about here]

Figure 3 shows that subjects grossly miscalculate the probability of casting a pivotal vote. Quantitatively, the subjects state a probability much lower than the theoretical and actual probabilities. Qualitatively, subjects' beliefs seem to be strongly affected by the distribution of preferences across teams. A low probability is attributed to situations

with a large difference in the number of supporters between the two teams, whereas the probability shows an important increase for closely divided teams. On average, the subjects stated a probability of 36.14% (with a standard deviation of 21.41%) of casting a pivotal vote when the difference between the teams is one. The stated probability decreases to 25.54% and 20.48% as the differences in team sizes increases to three and five respectively (the corresponding standard deviations are 17.53% and 18.01%). For teams of size seven the reported probability is 23.04% (the standard deviation is 27.08%). The subjects' estimates for a close distribution of preferences is significantly different than their estimates for the rest of the distributions (p < 0.001).

It follows from the previous subsection (see Figure 2) that not only subjects' beliefs of casting a pivotal vote are relatively higher for closely divided electorates, but also their propensity to vote increases for these electorates. As a consequence, the actual probability of casting a pivotal vote decreases in closely divided electorates. Hence, an increase in the subjects' beliefs of casting a pivotal vote brings about a decrease in the actual probability of being pivotal.

For a given distribution of preferences subjects, for the most part, attach a higher probability of casting a pivotal vote when they belong to the majority relative to the probability attached when subjects belong to the minority.¹¹ Contrary to our results from the previous section the differences are statistically significant for every distribution

 $^{^{10}}$ The probability stated for distributions with a division of two versus five is significantly different than that reported for distributions of one versus six (p < 0.001). Neither series, however, is significantly different from the probabilities reported for teams of size seven (p = 0.28 and 0.18, respectively).

¹¹In the only exception, subjects in the second survey assigned a higher probability of casting a pivotal vote to a team of size one (25.02%) than to a team of size six (21.48%).

of preferences. The p-value that subjects in the majority state a higher probability of casting a pivotal vote than subjects in the minority is below 0.04 for a distribution of four versus three subjects. This value decreases to 0.03 and to 0.003 as the difference between the teams increases to three and five, respectively.¹²

A comparison of the subjects' beliefs and their participation decisions leads us to conjecture that the subjects' behavior is at least partially accounted by their beliefs. Simply put, subjects may overestimate the likelihood of casting a pivotal vote and act rationally based on their inflated probability assessment. To test this hypothesis we estimate equation (4.1) replacing $Majority_{i,t}$, the explanatory variable that captured the relative effect of belonging to the majority, by each subject's beliefs of casting a pivotal vote conditional on the size of the subject's team.

The estimated coefficients appear on Column (2) of Table 2. The results are qualitatively similar to the ones observed in Column (1); that is, subjects' beliefs explain their behavior only when the distribution of preferences is closely divided between the two teams. Intuitively, in close elections subjects believe that there is a higher probability of casting a pivotal vote when they belong to the majority; these beliefs lead subjects to increase their relative frequency of voting when they indeed belong to the large group. Quantitatively, however, the coefficient for the subjects' beliefs is smaller than the coefficient estimated in Column (1). In particular, a ten percentage point increase in the belief of casting a pivotal vote when belonging to a team of size four causes a five percent

¹²All the conclusions above are reached using each subject's average of both surveys for a given team size. Interestingly, the subjects' responses to the surveys are not significantly different for any team size. Our results don't change if we use either survey instead of the subjects' average of the two surveys.

increase in the probability of voting of the average subject.

When the sizes of the teams aren't closely divided the subjects' beliefs don't play a significant role in their participation decision. In these situations, as was concluded before, the best predictor for a subject's participation in the second stage of a round is the subject's action in the first stage of the round.

The next section tests the external validity of our main experimental observation using data from gubernatorial elections in the U.S.

5. Evidence from Gubernatorial Elections in the US

This section's exercise is mainly intended to assess the external validity of our main experimental result. Using a newly culled data set on gubernatorial elections in the U.S. we test whether, in close elections, the difference in the actual vote tally between the party slightly leading according to the polls and the other party is larger than the one predicted by the polls. This hypothesis emanates directly from our experimental results and, to the best of our knowledge, has never been addressed in the vast extant empirical literature on voter turnout.

For the purposes of our empirical exercise we use as our proxy for closeness in the distribution of preferences the results of pre-election polls on gubernatorial races in the U.S., between 1990 and 2005. These polls, conducted by an independent polling firm (Mason-Dixon Polling and Research, Inc.), were published by newspaper media within

one week before the elections.¹³ The polls are supposed to be extremely accurate. They are published right before the elections and report results based only on likely voters. Therefore, the polls already incorporate other factors that affect participation (e.g. candidates' spending and mobilizations' effects). Hence, any systematic difference between the polls and the electoral results may be attributed, at least partially, to effects that the poll has on the electorate – effects that were not taken into account by the polling company.

Our data set consists of 143 gubernatorial elections in 47 states. These are all the elections between 1990 and 2005 where Mason-Dixon Polling and Research, Inc. conducted a public opinion poll within one week before the elections and where a third party didn't receive more than 30 percent of the election's votes.

The main two variables of interest are the differences in the vote share for the leading party minus the vote share for the trailing party according to the polls, and the electoral results. Let us denote by

$$DP = L_p - T_p$$

the difference in the vote share of the leading and trailing parties according to the polls, and denote by

$$DE = L_e - T_e$$

the corresponding difference between the two parties according to the electoral results.¹⁴

¹³According to Matsusaka and Palda (1993, p. 861) "the ideal measure would be survey predictions from opinion polls taken the day before the election." Our data come as close as possible to that ideal.

 $^{^{14}}$ For a given election the classification of the parties as leading or trailing is fixed. Therefore, DP is always positive whereas DE may be negative if the winner of the election is the party trailing in the published public opinion poll. This occurs for 11 observations in our sample.

Our exercise focuses on deviations of the electoral results from the polls predictions, DE - DP, and how these deviations correlate with the size of DP.

Table 3 presents summary statistics for the variables of interest. The table differentiates between elections where the difference in the support between the two parties according to the polls was less than 10 percentage points and the rest of the elections.¹⁵

[Table 3 about here]

The table clearly illustrates the main difference between closely divided electorates and the rest. The mean difference between DE and DP is positive for elections expected to be close and negative for the rest of the elections. Moreover, the mean deviation between the electoral results and the polls' predictions is higher, in absolute value, for closely divided electorates than for the rest of the electorates (0.0154 and -0.0071 percentage points, respectively) even though the latter group of elections has a higher variance.

Figure 4 depicts DE - DP on the vertical axis and DP on the horizontal axis. The figure includes the resulting curve according to the predicted value of DE - DP based on the estimation of a fractional polynomial of DP, along with the confidence interval of the mean (calculated using robust standard errors).

[Figure 4 about here]

 $^{^{15}}$ The chosen cutoff of 10 percent is the level of closeness that emerges endogenously from the analysis below.

The observed pattern in Figure 4 is consistent with our experimental results. Accordingly, for polls pointing to a narrow margin between the candidates we observe a bandwagon effect, whereby supporters of the leading candidate increase their participation relative to supporters of the trailing candidate. This is the case for polls predicting a difference smaller than 22 percentage points between the two parties. The bandwagon effect is particularly strong for DP values between 4 and 10 percentage points. In this range DE - DP is statistically greater than zero at the 2.5% significance level. The statistically greater than zero at the 2.5% significance level.

In electoral contests that are expected to be one-sided the above effect isn't present. For these contests the predicted value of DE - DP is decreasing as the difference in the support for the two parties according to the polls increases. As predicted by the theoretical models, free riding of supporters of the large party seems to be behind the negative slope of DE - DP in lopsided divided electorates.

Summing up, the evidence presented above is consistent with our experimental results. We need to stress that the empirical analysis is correlational in nature – it can't by itself establish causality. Combined with our experimental results, however, the empirical evidence strengthens the case for the existence of a bandwagon effect of public opinion polls in closely divided electorates.

¹⁶One may think that the publication of the poll may not only affect the voters' participation decision but their preferences as well. Since we restrict our attention to polls published within one week of the actual elections we believe that this effect isn't of an important magnitude.

 $^{^{17}}$ If we restrict the estimation to be linear DE - DP is statistically greater than zero at the 2.5% significance level for every DP value lower than 0.1.

6. A Behavioral Model

Our main experimental results, confirmed using data on gubernatorial elections, cannot be accounted by the traditional rational choice approach to turnout. The results therefore call for an alternative behavioral explanation.

An alternative specification of the voters' utility function may help explain part of the behavior observed in the laboratory. According to the traditional approach each voter's benefit and cost of participation aren't affected by whether the voter is in the losing or winning side of the contest. Several papers, mainly interested in models of sequential voting, modify the voters' utility function to take into account the fact that voters experience a desire to vote for the winner (on top of the costs and benefits explicit in rational choice models). This approach, while able to explain bandwagons, assumes that the very behavior we need to explain is good for the voters. Moreover, the approach doesn't account for the stark difference observed in the voters' behavior between elections expected to be close and the rest of the elections.

In this section we propose an alternative approach that relaxes Nash equilibrium but is consistent with the voters' beliefs and behavior observed in the laboratory. The Nash equilibrium concept requires that players optimize with respect to beliefs which are consistent with the actual strategies of players. As we have already pointed out no such combination can support the behavior we observe in our findings. Our experimental results may

¹⁸Borgers (2004) mentions this possibility in a simultaneous voting game similar to the one we analyze here. Callander (2004) models this possibility explicitly in a sequential voting game by adding a positive parameter to a voter's utility function when the voter supports the winning alternative. See Morton and Williams (1999) and Battaglini et al. (2005) for experimental studies of sequential voting games.

be consistent with a weaker notion of rationality. Are there "reasonable" beliefs that we can attribute to voters under which voters' best responses will be akin to the observed behavior (without these beliefs being consistent with the actual voters' strategies)?

Herein we present a set of reasonable beliefs which will satisfy these conditions: If voters believe that in a close election the probability of voting is similar for all the individuals and those probabilities are sufficiently high, then optimal behavior with respect to these beliefs gives rise to voting patterns consistent with the ones documented in the previous sections. Formally,

Proposition 1: Suppose individuals believe that voters in the majority vote with probability q and voters in the minority vote with probability r, with $|q-r| < \varepsilon$, for some sufficiently small $\varepsilon > 0$. If $r \ge 1/2$ the probability of casting a pivotal vote is higher for a voter in the majority than for a voter in the minority.

To illustrate the intuition behind the proposition let us consider the case of a closely divided electorate when individuals believe that r is close to one. In this case, an individual that supports the large team believes it is very likely that her vote may break a tie. On the contrary, an individual that supports the small team believes that her team will loose the election regardless of her choice. Thus, under the conditions of Proposition 1 individuals believe that there is a higher probability of casting a pivotal vote when they support the large team. This result is consistent with the beliefs stated by the subjects in their answers to the surveys.¹⁹

¹⁹The subjects' beliefs are similar to a "level-1" individual best responding to "level-0" individuals in the theoretical framework developed by Crawford and Iriberri (2005).

The main condition behind Proposition 1 is the individuals' beliefs that all voters mix with similar probabilities. Although these beliefs cannot be part of equilibrium with groups of unequal sizes, they seem reasonable when the preferences of the electorate are almost equally split between the two alternatives. This may explain why the bandwagon effect of polls occurs only when the electorate is closely divided.

7. Conclusions

This paper studies the effect that information on the voters' distribution of preferences has on turnout. The main finding is that the observed increase in turnout when the distribution of preferences is closely divided is heterogenous across teams of different sizes. In particular, the increase in turnout is significantly larger for the alternative with a slight majority according to the poll. That is, polls have a bandwagon effect whereby the frontrunner alternative increases its relative support in the elections. This effect, observed only in close elections, is not a consequence of voters changing their preferences. Rather, it is entirely driven by individuals that already supported the leading team voting with a relatively higher frequency.

We showed that the bandwagon effect of polls in closely divided electorates is a direct consequence of the subjects' beliefs. That is, for closely divided electorates we observe that subjects overestimate the probability of casting a pivotal vote and behave according to those beliefs. On the contrary, subjects' beliefs don't explain their actions in electorates that are lopsided divided. Rather, only subjects that voted with high frequencies regardless

of their beliefs or the distribution of preferences are the ones that participate in lopsided elections. This paper documented the bandwagon effect not only in the laboratory but also using data from U.S. gubernatorial elections in the last fifteen years.

It is noteworthy that the bandwagon effect cannot be accounted by the intuitions derived from theoretical models on the effect of public opinion polls on turnout. This theory is based on rational individuals holding the correct beliefs for every distribution of preferences. Hereby we proposed an alternative explanation consistent with the voters' beliefs and behavior observed in the laboratory. In particular, we presented a set of reasonable beliefs that can be attributed to voters under which utility maximization yields a behavioral pattern consistent with the bandwagon effect of polls in closely divided electorates.

Summing up, this paper discovered an anomalous behavioral pattern in the laboratory; it corroborated the external validity of this behavior for large electorates; and it presented an alternative rationale for the prevalence of bandwagon effects in close elections. Clearly, much work remains to be done for us to be able to understand what causes this effect. Currently, we are exploring the prevalence of this effect in general environments. It seems that subjects don't fully take into account information on an ex-post asymmetric distribution in environments that are ex-ante symmetric. This conjecture, if validated in the laboratory, has implications far beyond the context of voters' turnout.

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APPENDIX A

Definition of Totally Mixed Strategies Symmetric Bayesian Nash Equilibrium: Let us denote by n_A the number of voters that prefer alternative A and n_B the number of voters that prefer alternative B, with $n_A + n_B = n$. A totally mixed strategies symmetric BNE in this context corresponds to a probability of voting γ that satisfies the following two conditions:

$$\frac{2C}{V} = \left(\frac{1}{2}\right)^{n-1} \sum_{n_B=0}^{n-1} \binom{n-1}{n_B} \Pi(\gamma \mid A) \tag{A.1}$$

and

$$\frac{2C}{V} = \left(\frac{1}{2}\right)^{n-1} \sum_{n_A=0}^{n-1} \binom{n-1}{n_A} \Pi(\gamma | B)$$
 (A.2)

where

$$\Pi(\gamma \mid A) = \sum_{i=0}^{\min[n_A-1, n_B]} \binom{n_A-1}{i} \binom{n_B}{i} \gamma^{2i} (1-\gamma)^{n-1-2i} + \sum_{i=0}^{\min[n_A-1, n_B-1]} \binom{n_A-1}{i} \binom{n_B}{i+1} \gamma^{2i+1} (1-\gamma)^{n-2(i+1)}$$

denotes the probability that a voter that supports alternative A is pivotal when the rest of the voters vote with probability γ . Similarly,

$$\Pi(\gamma \mid B) = \sum_{i=0}^{\min[n_A, \; n_B-1]} \binom{n_A}{i} \binom{n_B-1}{i} \gamma^{2i} (1-\gamma)^{n-1-2i} + \sum_{i=0}^{\min[n_A-1, n_B-1]} \binom{n_A}{i+1} \binom{n_B-1}{i} \gamma^{2i+1} (1-\gamma)^{n-2(i+1)}$$

denotes the probability that a voter that supports alternative B is pivotal when the rest of the voters vote with probability γ . Thus, conditions (A.1) and (A.2) state that a voting

probability γ is a totally mixed symmetric BNE if and only if every individual is indifferent between voting and abstaining.

Definition of Totally Quasi-Symmetric Mixed Strategies Nash Equilibrium:

As defined by Palfrey and Rosenthal (1983, pp. 27), a pair of voting strategies (λ_A, λ_B) is a totally quasi-symmetric mixed strategy Nash equilibrium if and only if

$$\frac{2C}{V} = \sum_{i=0}^{\min[n_A - 1, n_B]} \binom{n_A - 1}{i} \binom{n_B}{i} \lambda_A^i (1 - \lambda_A)^{n_A - 1 - i} \lambda_B^i (1 - \lambda_B)^{n_B - i} + (A.3)$$

$$\sum_{i=0}^{\min[n_A - 1, n_B - 1]} \binom{n_A - 1}{i} \binom{n_B}{i + 1} \lambda_A^i (1 - \lambda_A)^{n_A - 1 - i} \lambda_B^{i+1} (1 - \lambda_B)^{n_B - 1 - i}$$

and

$$\frac{2C}{V} = \sum_{i=0}^{\min[n_A, n_B - 1]} {n_A \choose i} {n_B - 1 \choose i} \lambda_A^i (1 - \lambda_A)^{n_A - i} \lambda_B^i (1 - \lambda_B)^{n_B - 1 - i} + (A.4)^{n_A - i} \lambda_A^i (1 - \lambda_A)^{n_A - i} \lambda_B^i (1 - \lambda_B)^{n_B - 1 - i} + (A.4)^{n_A - 1} \sum_{i=0}^{\min[n_A - 1, n_B - 1]} {n_A - 1 \choose i + 1} \lambda_A^{i+1} (1 - \lambda_A)^{n_A - 1 - i} \lambda_B^i (1 - \lambda_B)^{n_B - 1 - i}$$

where the right hand side of (A.3) is simply the probability that a voter supporting alternative A is pivotal when the rest of the voters supporting A vote with probability λ_A and all the voters supporting B vote with probability λ_B . Similarly, the right hand side of (A.4) is the probability that a voter supporting alternative B is pivotal when all the voters supporting A vote with probability λ_A and the rest of the voters supporting B vote with probability λ_B . Therefore, (A.3) states a sufficient and necessary condition for λ_A to be a best response to λ_B and (A.4) states a sufficient and necessary condition for λ_B to

be a best response to λ_A .

Proof of Proposition 1: Let us assume that there are n voters, with $n \ge 3$ and odd. Let us say that n_A of the voters prefer alternative A and n_B of them prefer alternative B, with $n_A + n_B = n$ and $n_A < n_B$. Assume first that r = q. The probability that an individual that prefers alternative A is pivotal equals

$$\sum_{i=0}^{n_A-1} {n_A-1 \choose i} {n_B \choose i} q^{2i} (1-q)^{n-1-2i} + \sum_{i=0}^{n_A-1} {n_A-1 \choose i} {n_B \choose i+1} q^{2i+1} (1-q)^{n-2(i+1)}$$
 (A.5)

where the first term is the probability of observing a tie and the second term is the probability that alternative A loses the election by one vote. Similarly, the probability that an individual that prefers alternative B is pivotal equals

$$\sum_{i=0}^{n_A} \binom{n_A}{i} \binom{n_B-1}{i} q^{2i} (1-q)^{n-1-2i} + \sum_{i=0}^{n_A-1} \binom{n_A}{i+1} \binom{n_B-1}{i} q^{2i+1} (1-q)^{n-2(i+1)}. \quad (A.6)$$

Therefore, rearranging terms we can express the probability that an individual supporting B will break a tie minus the probability that an individual supporting A will break a tie as

$$\sum_{i=0}^{n_A-1} {n_B-1 \choose i} {n_A-1 \choose i} \frac{i(n_B-n_A)}{(n_B-i)(n_A-i)} q^{2i} (1-q)^{n-1-2i} + {n_B-1 \choose n_A} q^{2n_A} (1-q)^{n-1-2n_A}.$$
(A.7)

Similarly, subtracting from the second term of (A.5) the second term of (A.6) we obtain

$$\sum_{i=0}^{n_A-1} {n_B-1 \choose i+1} {n_A-1 \choose i} \frac{(n_B-n_A)}{(n_B-1-i)} q^{2i+1} (1-q)^{n-2(i+1)}. \tag{A.8}$$

Thus, the probability of casting a pivotal vote is greater when an individual supports the majority group B if and only if (A.7) is greater than (A.8); this is equivalent to

$$\sum_{i=0}^{n_A-1} {n_B-1 \choose i+1} {n_A-1 \choose i} \frac{(n_B-n_A)}{(n_B-1-i)} q^{2i+1} (1-q)^{n-1-2(i+1)} \left[q-(1-q)\right] > 0.$$

This inequality is satisfied if, and only if, q > 1/2. The more general result for $r \neq q$ with $|q - r| < \varepsilon$, for some $\varepsilon > 0$ follows immediately from the fact that the probability of casting a pivotal vote is continuous on q and r.

APPENDIX B

Experiment in Decision-Making

This is an experiment in decision-making. During the experiment, you will make decisions and the other participants will do so as well. Your decisions and the others' will determine the payment that you will receive according to rules that we will explain later on.

You will be paid in cash at the end of the experiment, exactly as the rules say.

Twenty-one people are participating in the experiment. The experiment will be conducted by means of computers. All decisions that you make during the experiment will be implemented by keying appropriate commands.

Please remain totally silent during the experiment and do not speak with the other participants. If you have a question of any kind, raise your hand and one of the supervisors will come over to you.

The experiment is composed of several rounds. Each round has two phases—Phase 1 and Phase 2.

At the beginning of the experiment, the twenty-one participants will be divided up randomly, by the computer, into three groups of seven participants apiece. You will not know whom among the other participants in the room belongs to your group. Each group plays among its own members only and is independent of the other groups in the room.

At the beginning of the experiment, each participant receives a participation fee of 80 tokens.

At the beginning of each round, each participant receives a message from the computer about whether the participant prefers the choice of the blue color or of the green color. The computer performs a separate draw for each participant as to his or her preference of the color to be chosen. Each color, green or blue, has a 50 percent probability of being chosen.

After the computer tells you which color you prefer—blue or green—you will be asked to make decisions:

- a. In Phase 1 of the game, you will have to decide whether to vote for the color that you prefer. For your decision to vote for the color that you prefer, you will pay four tokens. In Phase 1 of the game, you will be given no information about the color preferences of the other members of your group. After you make a decision, the computer will immediately move on to Phase 2 of the game.
- b. In Phase 2 of the game, you will receive information about the color preferences of the members of your group. For example, you will be told that two players in your group prefer to see the blue color chosen and that five members of your group prefer to see green chosen. Then you will be asked to decide whether to vote for the color that you prefer. For your decision to vote for the color that you prefer, you will pay four tokens.

At the end of Phase 2 of the game, the computer will collect the seven decisions of the members of your group and determine which color was chosen in Phase 1 of the game and which color was chosen in Phase 2 of the game. The color chosen is the one for which a majority of the group members votes.

If the color that you preferred was chosen in Phase 1, you will receive ten tokens, and the other members of the group who preferred this color will also receive ten tokens—whether they voted for it or not.

If the color that you preferred was chosen in Phase 2, you will receive ten tokens, and the other members of the group who preferred this color will also receive ten tokens—whether they voted for it or not.

Note:

The fact that many members of the group prefer a certain color does not mean that this color will be chosen. For a color to be chosen, it must receive an actual majority of votes.

If all seven members of the group decide not to vote for the colors that they prefer, the computer will select at random the color of choice in the respective phase of the game. The computer will do the same in the case of a **tie vote** among group members between green and blue. Each color, green or blue, has a 50 percent probability of being chosen.

After each round, each participant will receive feedback about:

- * the color chosen in Phase 1 of the round and the color chosen in Phase 2 of the round.
- * the number of tokens available to you at the beginning of the round.
- * the number of tokens that you paid out during the round (in return for the choice of voting for the color that you prefer in Phase 1 of the round and in Phase 2 of the round).
- * the number of tokens that you earned during the round (by having your color of preference chosen in Phase 1 of the round and Phase 2 of the round).
- * the number of tokens that you have accumulated thus far. The tokens that you accumulate by the end of the round will be available to you at the beginning of the next round.

After you receive the feedback, the next round begins. It, too, is composed of two phases, and in each phase you will be asked to decide whether you wish to vote for the color that you prefer.

Note: at the beginning of the next round, the groups will be recomposed. The computer will again divide the twenty-one participants in the room, at random, into three different groups of seven participants apiece. Again the computer will give you a message about the color that you prefer in the new round.

At the end of the experiment, the computer will add up the total tokens that you accumulated in all the rounds in the experiment. You will be paid in cash at the exchange rate of four tokens = NIS 1.

Examples: please fill in your answers in the appropriate boxes. If you have questions, raise your hand and the supervisor will come over to you.

Participant and his/her color	1-prefers green	2-prefers blue	3-prefers blue	4-prefers green	5-prefers blue	6-prefers green	7-prefers blue
preference							
Initial grant of	80	80	80	80	80	80	80
tokens at beginning							
of experiment							
Round 1							
Dl 1 .							
Phase 1:	Decided	Decided to	Decided	Decided to	Decided	Decided	Decided
decision-making without	to vote for	vote for	not to	vote for	not to	not to	not to
information about	preferred	preferred	vote for	preferred	vote for	vote for	vote for
other group	color	color	preferred	color	preferred	preferred	preferred
members'	COIOI	COIOI	color	COIOI	color	color	color
preferences			COIOI		COIOI	Color	COIOI
Which color		N	umber of vote	s in favor of g	reen:	<u> </u>	
received a majority							
of group members'		N	umber of vote	es in favor of	blue:		
votes in Phase 1?							
Round 1							
Phase 2:							
decision-making	Decided	Decided to	Decided to	Decided to	Decided	Decided	Decided to
with information	not to	vote for	vote for	vote for	not to	not to	vote for
about other group	vote for	preferred	preferred	preferred	vote for	vote for	preferred
members'	preferred color	color	color	color	preferred color	preferred color	color
preferences Which color	COIOI	N.	umber of vote	l s in favor of g			
received a majority		111	uniber of vote	s iii iavoi oi g	;ieeii		
of group members'		N	lumber of vote	es in favor of	hlue:		
votes in Phase 2?		1	umber of vou	es in favor of	oruc		
Total tokens paid							
by participant in							
both phases of the							
game for having							
chosen to vote for							
the color that							
he/she prefers							
Number of tokens							
earned by							
participant in this							
game for having							
chosen the color							
that he/she prefers Total tokens							
accumulated by							
participant thus far							
pai ucipant mus far							

<u>Figure 1</u>: Average Turnout Rate by Distribution of Preferences

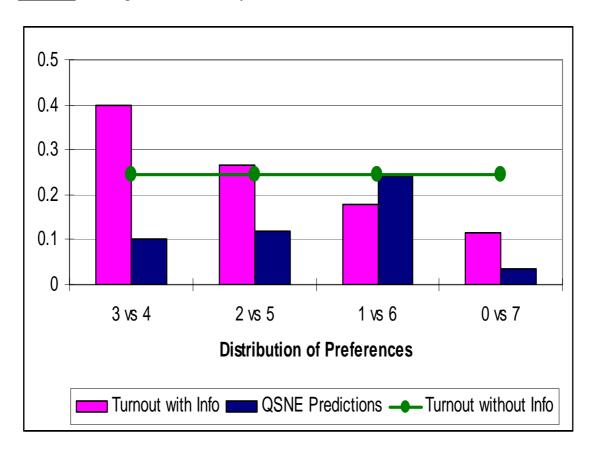


Figure 2: Average Turnout Rate by Team Size

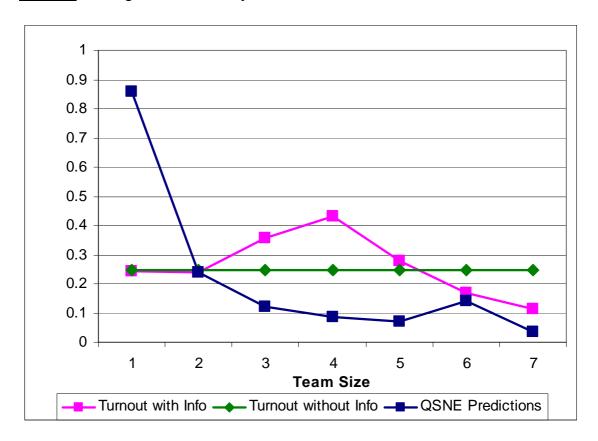
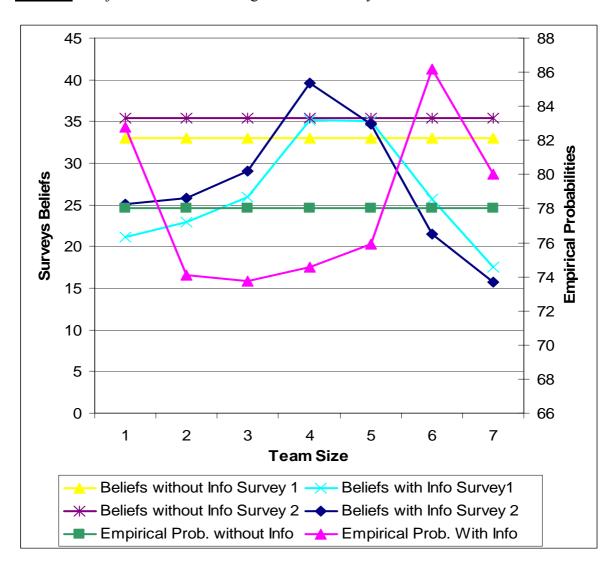
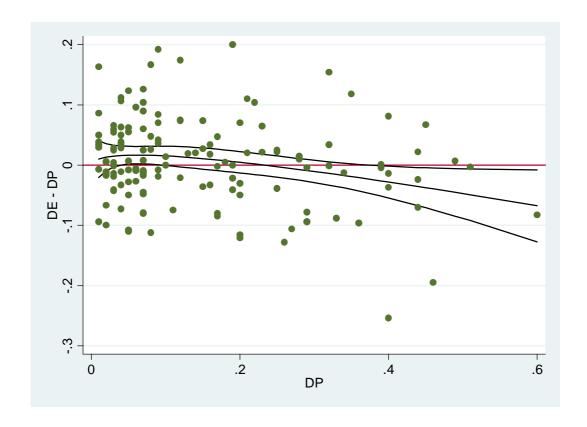


Figure 3: Subjects' Beliefs of Casting a Pivotal Vote by Team Size



<u>Figure 4</u>: Deviation of Electoral Results (DE) from Polls Predictions (DP) for Gubernatorial Elections, 1990 – 2005.



<u>**Table 1:**</u> Theoretical Predictions for the Two Different Frameworks with Benefits of Winning Equal to 10 and Costs of Voting Equal to 4.

Distribution of Preferences	3 vs. 4	2 vs. 5	1 vs. 6	0 vs. 7
Probability that a supporter of the large team votes according to the unique QSNE	0.0873	0.070805	0.13988	0.036508
Probability that a supporter of the small team votes according to the unique QSNE	0.1229	0.24001	0.86012	
Probability of voting according to the unique BNE	0.0807			

<u>Table 2</u>: Random Effect Probit Estimates of Voting Decisions in the Second Stage (Standard errors in Parentheses).

		(1)			(2)	
Distribution of Preferences	3 vs. 4	2 vs. 5	1 vs. 6	3 vs. 4	2 vs. 5	1 vs. 6
Constant	-0.522***	-1.039***	-1.030***	-0.740***	-0.977***	-1.446***
	(0.170)	(0.243)	(0.407)	(0.200)	(0.240)	(0.359)
First Stage Voting Decision	0.207	0.342^{*}	0.921***	0.232	0.360^{*}	0.954***
	(0.143)	(0.213)	(0.324)	(0.145)	(0.212)	(0.335)
Round	-0.020***	-0.002	-0.002	-0.020***	-0.002	-0.004
	(0.009)	(0.015)	(0.022)	(0.009)	(0.014)	(0.023)
Majority	0.305***	0.175	-0.339			
	(0.108)	(0.184)	(0.332)			
Pivotal Beliefs				0.012***	0.002	0.005
				(0.004)	(0.004)	(0.005)
Observations	854	378	203	854	378	203
Number of Subjects	84	84	72	84	84	72

^{**, **, *} Indicates Statistical Significance at the 1, 5 and 10 percent level respectively.

<u>Table 3</u>: Average Difference in Vote Share between Leading and Trailing Party according to Polls and Elections, 1990 – 2005 (Standard errors in Parentheses)

	(1) Elections Expected to be Close (DP < 0.10)	(2) Elections Not Expected to be Close (DP \geq 0.10)
Difference in Vote Share between Leading and Trailing Party according to Polls (DP)	0.0491 (0.0258)	0.2663 (0.1147)
Difference in Vote Share between Leading and Trailing Party according to Elections (DE)	0.0645 (0.0714)	0.2592 (0.1235)
Number of Observations	78	65

Sources: Polls' data obtained from Mason-Dixon Polling and Research, Inc.